Some directions in Perfect Simulation

Laird Breyer, mailto:laird@lbreyer.com

joint work with Gareth Roberts (Lancaster) The work presented here has involved the following institutions:

University of Lancaster

Aalborg University

Ж

Università di Roma Tre



In this talk, I shall present loosely some recent ideas related to Perfect Simulation, including:

In this talk, I shall present loosely some recent ideas related to Perfect Simulation, including:

• what Read-Once CFTP means

In this talk, I shall present loosely some recent ideas related to Perfect Simulation, including:

- what Read-Once CFTP means
- some old and new coupling constructions

In this talk, I shall present loosely some recent ideas related to Perfect Simulation, including:

- what Read-Once CFTP means
- some old and new coupling constructions
- an example of perfect simulation

In this talk, I shall present loosely some recent ideas related to Perfect Simulation, including:

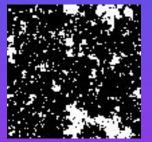
- what Read-Once CFTP means
- some old and new coupling constructions
- an example of perfect simulation

PS is still at the beginning, and new methods are discovered regularly...

In MCMC, we commonly have two distinct goals:

In MCMC, we commonly have two distinct goals:

Simulation



Integration

 $\mathbb{P}(A \mid B) = \int_{A} \mathbb{P}(B \mid \omega) \mathbb{P}(d\omega) / \mathbb{P}(B)$

In MCMC, we commonly have two distinct goals:

Simulation

Integration

 $\mathbb{P}(A \mid B) = \int_{A} \mathbb{P}(B \mid \omega) \mathbb{P}(d\omega) / \mathbb{P}(B)$

PS solves the simulation problem completely,

In MCMC, we commonly have two distinct goals:

Simulation

Integration

$$\mathbb{P}(A \mid B) = \int_{A} \mathbb{P}(B \mid \omega) \mathbb{P}(d\omega) / \mathbb{P}(B)$$

PS solves the simulation problem completely, and reduces integration problems to Classical Statistics.

Building Markov chains

There are several equivalent definitions for Markov chains. For PS, we use a computational definition:

A sequence X_1, X_2, \ldots is a Markov chain if there exists an IID series of random functions $F_1(x), F_2(x), \ldots$ such that

 $F_1(X_1) = X_2, \quad F_2(X_2) = X_3, \quad \dots, \quad F_k(X_k) = X_{k+1}$

Building Markov chains

There are several equivalent definitions for Markov chains. For PS, we use a computational definition:

A sequence $X_1, X_2,...$ is a Markov chain if there exists an IID series of random functions $F_1(x), F_2(x),...$ such that

> $F_1(X_1) = X_2, \quad F_2(X_2) = X_3, \quad \dots, \quad F_k(X_k) = X_{k+1}$ X_{s+12} F_{s+2} F_{s+3} F_{s+4} F_{s+5} F_{s+6} F_{s+7} F_{s+8} F_{s+9} F_{s+10} F_{s+11} F_{s+12}

The stationary density

Consider the chain over a larger time scale: the stationary density π satisfies $\pi P = \pi$, where

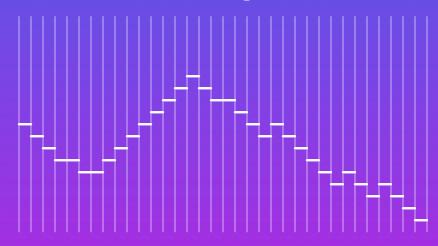
 $P(x, dy) = \mathbb{P}(F_t(x) \in dy)$ for all t.

The stationary density

Consider the chain over a larger time scale: the stationary density π satisfies $\pi P = \pi$, where

 $P(x, dy) = \mathbb{P}(F_t(x) \in dy)$ for all t.

Here is a path:

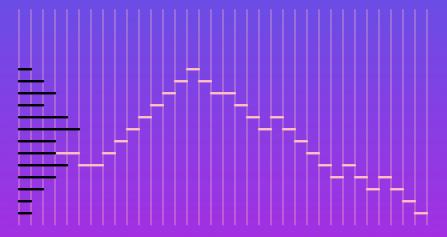


The stationary density

Consider the chain over a larger time scale: the stationary density π satisfies $\pi P = \pi$, where

 $P(x, dy) = \mathbb{P}(F_t(x) \in dy)$ for all t.

After reordering the occupied states, we get:

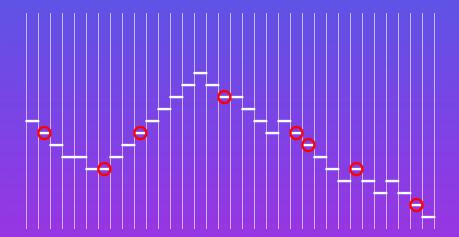


Read Once Perfect Simulation

Wilson (1999) showed how to recognize random times T_2, T_3, \ldots when the chain is *exactly* in equilibrium:

Read Once Perfect Simulation

Wilson (1999) showed how to recognize random times T_2, T_3, \ldots when the chain is *exactly* in equilibrium:



Read Once Perfect Simulation

Wilson (1999) showed how to recognize random times T_2, T_3, \ldots when the chain is *exactly* in equilibrium:



How to find the perfect samples

- Trace all possible paths under the mappings F_1 , F_2 ,..., looking for *coalescence* within m steps:
- Let $G_1(x)$, $G_2(x)$,... be defined for all x by:

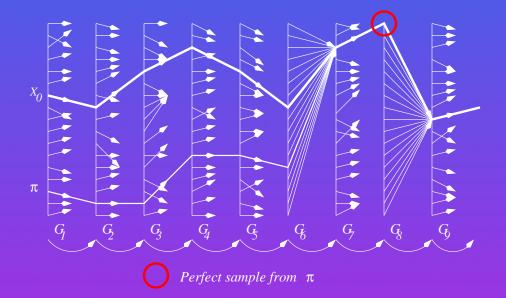
$$G_1(x) = F_m(F_{m-1}(\cdots F_1(x)\cdots))$$

$$G_2(x) = F_{2m}(F_{2m-1}(\cdots F_{m+1}(x)\cdots))$$

and set

 $T_{1} = \min\{u \ge 0 : G_{u}(x) = G_{u}(y) \text{ for all } x, y\}$ $T_{k+1} = \min\{u \ge T_{k} : G_{u}(x) = G_{u}(y) \text{ for all } x, y\}$

How to find the perfect samples



Let $\epsilon = \mathbb{P}(T_1 \leq 1)$ and write

 $\overline{\mathbb{P}(G_1(x) \in dy)} = (1 - \epsilon)\mathbb{P}(G_1(x) \in dy \mid T_1 > 1) + \epsilon\mathbb{P}(G_1(x) \in dy \mid T_1 \le 1)$ $= (1 - \epsilon)Q(x, dy) + \epsilon\mu(dy).$

Then $\pi = \epsilon \sum_{s=0}^{\infty} (1-\epsilon)^s \mu Q^s$.

Let $\epsilon = \mathbb{P}(T_1 \leq 1)$ and write

 $\overline{\mathbb{P}(G_1(x) \in dy)} = (1 - \epsilon)\mathbb{P}(G_1(x) \in dy \mid T_1 > 1) + \epsilon\mathbb{P}(G_1(x) \in dy \mid T_1 \le 1)$ $= (1 - \epsilon)Q(x, dy) + \epsilon\mu(dy).$

Then $\pi = \epsilon \sum_{s=0}^{\infty} (1 - \epsilon)^s \mu Q^s$. *Proof.* Using stationarity $\pi P = \pi$, we have

 $(1-\epsilon)^k \pi Q^k = (1-\epsilon)^{k-1} \pi (P-\epsilon\mu) Q^{k-1}$

Let $\epsilon = \mathbb{P}(T_1 \leq 1)$ and write

$$\begin{split} \mathbb{P}(G_1(x) \in dy) &= (1-\epsilon) \mathbb{P}(G_1(x) \in dy \mid T_1 > 1) + \epsilon \mathbb{P}(G_1(x) \in dy \mid T_1 \le 1) \\ &= (1-\epsilon) Q(x, dy) + \epsilon \mu(dy). \end{split}$$

Then $\pi = \epsilon \sum_{s=0}^{\infty} (1 - \epsilon)^s \mu Q^s$. *Proof.* Using stationarity $\pi P = \pi$, we have

$$(1-\epsilon)^k \pi Q^k = (1-\epsilon)^{k-1} \pi (P-\epsilon\mu) Q^{k-1}$$

= $(1-\epsilon)^{k-1} \pi Q^{k-1} - \epsilon (1-\epsilon)^{k-1} \mu Q^{k-1}$

Let $\epsilon = \mathbb{P}(T_1 \leq 1)$ and write

 $\mathbb{P}(G_1(x) \in dy) = (1 - \epsilon)\mathbb{P}(G_1(x) \in dy \mid T_1 > 1) + \epsilon\mathbb{P}(G_1(x) \in dy \mid T_1 \le 1)$ $= (1 - \epsilon)Q(x, dy) + \epsilon\mu(dy).$

Then $\pi = \epsilon \sum_{s=0}^{\infty} (1 - \epsilon)^s \mu Q^s$. *Proof.* Using stationarity $\pi P = \pi$, we have

$$(1-\epsilon)^{k}\pi Q^{k} = (1-\epsilon)^{k-1}\pi (P-\epsilon\mu)Q^{k-1}$$

= $(1-\epsilon)^{k-1}\pi Q^{k-1} - \epsilon(1-\epsilon)^{k-1}\mu Q^{k-1}$
= \cdots =
= $\pi - \epsilon \sum_{s=1}^{k} (1-\epsilon)^{k-s}\mu Q^{k-s}.$

Let $\epsilon = \mathbb{P}(T_1 \leq 1)$ and write

 $\mathbb{P}(G_1(x) \in dy) = (1 - \epsilon)\mathbb{P}(G_1(x) \in dy \mid T_1 > 1) + \epsilon\mathbb{P}(G_1(x) \in dy \mid T_1 \le 1)$ $= (1 - \epsilon)Q(x, dy) + \epsilon\mu(dy).$

Then $\pi = \epsilon \sum_{s=0}^{\infty} (1 - \epsilon)^s \mu Q^s$. *Proof.* Using stationarity $\pi P = \pi$, we have

$$(1-\epsilon)^{k}\pi Q^{k} = (1-\epsilon)^{k-1}\pi (P-\epsilon\mu)Q^{k-1}$$

= $(1-\epsilon)^{k-1}\pi Q^{k-1} - \epsilon(1-\epsilon)^{k-1}\mu Q^{k-1}$
= \cdots =
= $\pi - \epsilon \sum_{s=1}^{k} (1-\epsilon)^{k-s}\mu Q^{k-s}.$

Let $k \to \infty$, then left side goes to zero.

How to ensure coupling occurs

Again assume that we have an IID sequence F_1, F_2, \ldots of random maps satisfying

$$P(x, dy) = \mathbb{P}(F_t(x) \in dy) \quad \text{for all } t.$$
(1)

With $F_t(x)$ arbitrary, coalescence of Markov chain sample paths need not ever happen. We must modify the maps subject to preserving the condition (1).

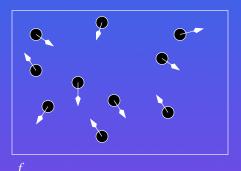
• Normally at each iteration, we use a randomly generated function $F_t(x)$ such that $X_{t+1} = F_t(X_t)$.

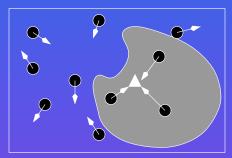
- Normally at each iteration, we use a randomly generated function $F_t(x)$ such that $X_{t+1} = F_t(X_t)$.
- If $F_t(x) = F_t(x')$ holds for some x, x', then there is the *possibility* of coupling.

- Normally at each iteration, we use a randomly generated function $F_t(x)$ such that $X_{t+1} = F_t(X_t)$.
- If $F_t(x) = F_t(x')$ holds for some x, x', then there is the *possibility* of coupling.
- Strategy: From F, generate a new random function $C_Y(F)$ which has a higher chance of coupling.

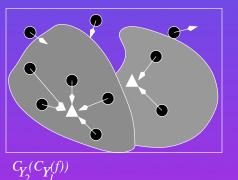
- Normally at each iteration, we use a randomly generated function $F_t(x)$ such that $X_{t+1} = F_t(X_t)$.
- If $F_t(x) = F_t(x')$ holds for some x, x', then there is the *possibility* of coupling.
- Strategy: From F, generate a new random function $C_Y(F)$ which has a higher chance of coupling.
- **Definition:** Let Y be independent of F, with $\mathbb{P}(Y = y) = q(y)$ say.

$$\mathcal{C}_{Y}(F)(x) = \begin{cases} Y & \text{if } \frac{p(x,Y)q(F(x))}{p(x,F(x))q(Y)} > U[0,1] \\ F(x) & \text{otherwise.} \end{cases}$$



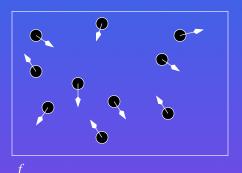


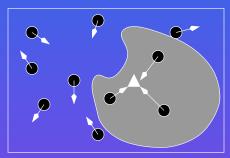




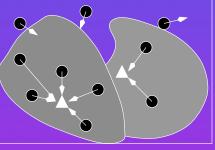
 \square

 $\overline{C_{Y_5}(C_{Y_4}(C_{Y_2}(C_{Y_2}(C_{Y_1}(f)))))}$

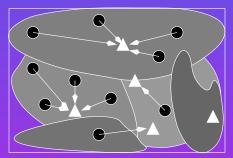












 $C_{Y_5}(C_{Y_4}(C_{Y_3}(C_{Y_2}(C_{Y_1}(f))))))$

We call the triangles catalysts. Click here.

Why it works

• In the formula

$$\mathcal{C}_{Y}(F)(x) = \begin{cases} Y & \text{if } \frac{p(x,Y)q(F(x))}{p(x,F(x))q(Y)} > U[0,1] \\ F(x) & \text{otherwise.} \end{cases}$$

fix x and define $\Pi(y) = p(x, y)$, X = F(x) and $X' = \mathcal{C}_Y(F)(x)$.

Why it works

• In the formula

$$\mathcal{C}_{Y}(F)(x) = \begin{cases} Y & \text{if } \frac{p(x,Y)q(F(x))}{p(x,F(x))q(Y)} > U[0,1] \\ F(x) & \text{otherwise.} \end{cases}$$

fix x and define $\Pi(y) = p(x, y)$, X = F(x) and $X' = C_Y(F)(x)$.

• We get

$$X' = \begin{cases} Y & \text{if } \frac{\Pi(Y)q(X)}{\Pi(X)q(Y)} > U[0,1] \\ X & \text{otherwise} \end{cases}$$

Why it works

• In the formula

$$\mathcal{C}_{Y}(F)(x) = \begin{cases} Y & \text{if } \frac{p(x,Y)q(F(x))}{p(x,F(x))q(Y)} > U[0,1] \\ F(x) & \text{otherwise.} \end{cases}$$

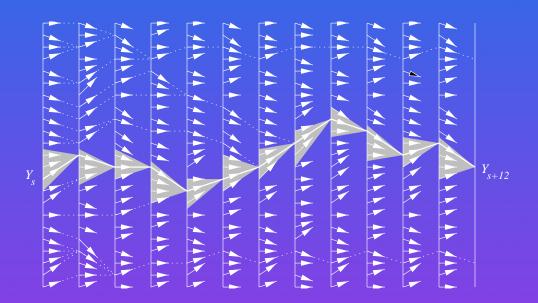
fix x and define $\Pi(y) = p(x, y)$, X = F(x) and $X' = C_Y(F)(x)$.

• We get

$$X' = \begin{cases} Y & \text{if } \frac{\Pi(Y)q(X)}{\Pi(X)q(Y)} > U[0,1] \\ X & \text{otherwise} \end{cases}$$

• Therefore, since $\Pi(dy) = \mathbb{P}(X \in dy)$ by assumption, it follows that $\mathbb{P}(X' \in dy) = \Pi(dy)$ also.

Extensions



- At every update, let the catalyst Y_t come from an independent Markov chain with identical transition probabilities. Great for Gibbs samplers!
- We have the formula

$$\|\mathbb{P}(X_t \in \cdot) - \pi\|_{\mathrm{TV}} \le \mathbb{E} \prod_{s=0}^t \left(1 - \frac{p(X_s, Y_{s+1})p(Y_s, X_{s+1})}{p(X_s, X_{s+1})p(Y_s, Y_{s+1})} \right)_+,$$

Pump Example (Autogamma)

• Let $x = (\beta, \lambda_1, \dots, \lambda_{10})$, simulate

$$\pi(x) = \exp\left\{ (10\alpha + \gamma - 1) \log \beta - \delta\beta + \sum_{k=1}^{10} \left((s_k + \alpha - 1) \log \lambda_k - (\beta + t_k)\lambda_k \right) \right\},$$

Pump Example (Autogamma)

• Let
$$x = (\beta, \lambda_1, \dots, \lambda_{10})$$
, simulate

$$\pi(x) = \exp\left\{ (10\alpha + \gamma - 1) \log \beta - \delta\beta + \sum_{k=1}^{10} \left((s_k + \alpha - 1) \log \lambda_k - (\beta + t_k)\lambda_k \right) \right\},\$$

• Gibbs sampler: One sweep is $f : (\beta, \lambda_1, \dots, \lambda_{10}) \mapsto (\beta', \lambda'_1, \dots, \lambda'_{10})$, where

$$\beta' \sim \pi_0(\cdot \mid \lambda_1, \dots, \lambda_{10}) = \Gamma(\gamma + 10\alpha, \delta + \sum_{k=1}^{10} \lambda_k),$$
$$\lambda'_k \sim \pi_k(\cdot \mid \beta') = \Gamma(\alpha + s_k, \beta' + t_k), \quad k = 1, \dots, 10$$

• Transition density is

 $p(\beta, \lambda_1, \dots, \lambda_{10}; b, l_1, \dots, l_{10}) = \pi_0(b \mid \lambda_1, \dots, \lambda_{10}) \prod_{k=1}^{10} \pi_k(l_k \mid b).$

• Transition density is

 $p(\beta, \lambda_1, \dots, \lambda_{10}; b, l_1, \dots, l_{10}) = \pi_0(b \mid \lambda_1, \dots, \lambda_{10}) \prod_{k=1}^{10} \pi_k(l_k \mid b).$

• Take $Y \sim q$ where

$$q(y_0, y_1, \dots, y_{10}) = \pi_0(y_0 \mid \lambda_1^*, \dots, \lambda_{10}^*) \prod_{k=1}^{10} \pi_k(y_k \mid y_0),$$

• After simplification,

$$\mathcal{C}_{Y}(f)(x) = egin{cases} Y & ext{if e} \ (eta', \lambda'_{1}, \dots, \lambda'_{10}) & ext{othermality} \end{cases}$$

if
$$\exp\left((\beta' - Y_0)(|\lambda| - |\lambda^*|)\right) > \xi$$
, otherwise.

• After simplification,

$$\mathcal{C}_{Y}(f)(x) = \begin{cases} Y & \text{if } \exp\Big((\beta' - Y_0)(|\lambda| - |\lambda^*|)\Big) > \xi, \\ (\beta', \lambda'_1, \dots, \lambda'_{10}) & \text{otherwise.} \end{cases}$$

• Thus

$$\operatorname{Basin}(Y, f, \xi) = \Big\{ x : \Big(\psi_0 - (\delta + |\lambda|) Y_0 \Big) \Big(|\lambda| - |\lambda^*| \Big) > \log \xi \Big\},\$$

• After simplification,

$$\operatorname{Basin}(Y, f, \xi) = \left\{ x : |\lambda| \in \left[\frac{-b - \sqrt{b^2 - 4ac}}{2a}, \frac{-b + \sqrt{b^2 - 4ac}}{2a} \right] \right\},$$

where $a = Y_0$, $b = \log \xi - Y_0(|\lambda^*| - \delta)$ and $c = \delta(\log \xi - Y_0 |\lambda^*|) - \psi_0$.

Conclusions

• If we can recognize the basins, we can also recognize when the maps G_u coalesce.

Conclusions

- If we can recognize the basins, we can also recognize when the maps G_u coalesce.
- If we can recognize map coalescence, we can also do Perfect Simulation.

Conclusions

- If we can recognize the basins, we can also recognize when the maps G_u coalesce.
- If we can recognize map coalescence, we can also do Perfect Simulation.
- You can find a simulation by clicking here.

Learning more

Preprints and tutorials may be found at the following two sites:

- http://www.dimacs.rutgers.edu/~dbwilson/exact.html
- http://www.lbreyer.com